

Central Limit Theorems for Percolation Models

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Let $p \neq 1/2$ be the open-bond probability in Broadbent and Hammersley's percolation model on the square lattice. Let W_x be the cluster of sites connected to x by open paths, and let $\{\gamma(n)\}$ be any sequence of circuits with interiors $|\dot{\gamma}(n)| \rightarrow \infty$. It is shown that for certain sequences of functions $\{f_n\}$, $S_n = \sum_{x \in \dot{\gamma}(n)} f_n(W_x)$ converges in distribution to the standard normal law when properly normalized. This result answers a problem posed by Kunz and Souillard, proving that the number S_n of sites inside $\gamma(n)$ which are connected by open paths to $\gamma(n)$ is approximately normal for large circuits $\gamma(n)$.

KEY WORDS: Percolation; asymptotic normality; circuits; semi-invariants.

1. INTRODUCTION

The percolation model of Broadbent and Hammersley⁽²⁾ has received considerable recent attention in both the mathematics and physics literature (see Refs. 5–8, 10, 14–16). In this paper we will verify a conjecture of Kunz and Souillard (Section 6 in Ref. 10) and prove a general central limit theorem.

We will consider only the bond percolation model on the square lattice, although some of our methods should work for other models. Let \mathbb{Z}_2 be the set of points in the plane with integer coefficients, and for $x, y \in \mathbb{Z}_2$, $x = (x_1, x_2), y = (y_1, y_2)$, write $d(x, y) = |x_1 - y_1| + |x_2 - y_2|$. The points of \mathbb{Z}_2 will be called *sites*, and the line segments joining sites x and y with $d(x, y) = 1$ will be called *bonds*. The origin is the site $(0, 0)$, denoted by 0 . Each bond is declared *open* with probability p or *closed* with probability $1 - p$ independently of all other bonds. We will assume throughout that $0 < p < 1$.

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A *path* from x to y is an alternating sequence of sites and bonds of the form $x_0, v_1, x_1, v_2, \dots, v_n, x_n$, where $x_0 = x, x_n = y, d(x_i, x_{i+1}) = 1$, and v_i is the bond joining x_{i-1} and x_i . A *circuit* γ is a path $x_0, v_1, x_1, \dots, v_n, x_n$ such that x_0, x_1, \dots, x_{n-1} are distinct and $x_0 = x_n$. Thus a circuit is a special type of a simple closed curve. Let $\dot{\gamma}$ be the set of sites strictly *inside* γ and let $\partial\dot{\gamma}$ be the *inside boundary*, $\partial\dot{\gamma} = \{x \in \dot{\gamma} \mid \text{there exists } y \in \gamma \text{ with } d(x, y) = 1\}$. The notation $x \sim y$ means there is a path from x to y with *all bonds open*, and for any $A \subset \mathbb{Z}_2, x \sim A$ means $x \sim y$ for some $y \in A$. The number of sites in A is denoted $|A|$.

The *open cluster* W_x at x is defined to be the set of all sites y such that $x \sim y$. If the four bonds of x are all closed, W_x is empty. For any $A \subset \mathbb{Z}_2$ let $W_A = \cup_{x \in A} W_x$. The usual interpretation of W_x is that it represents the set of sites which are “wetted” by placing a fluid source at x and allowing fluid to flow only along open bonds. It is now known (see Ref. 8) that $P(|W_x| = \infty)$ is zero for $p \leq 1/2$ and strictly positive for $p > 1/2$.

Given a circuit γ we can alter our viewpoint by putting fluid sources at each site of γ and asking how many sites *inside* γ are “wetted.” In Ref. 10 Kunz and Souillard conjectured that the number of such sites should be approximately normal for large γ . We verify this conjecture with the following theorem.

Theorem 1. Assume $p \neq 1/2$ and define $\xi_\gamma(x) = 1_{\{W_x \cap \dot{\gamma} \neq \emptyset\}}$ and $S_\gamma = \sum_{x \in \dot{\gamma}} \xi_\gamma(x)$. Then there are finite, nonzero constants $c_1(p), c_2(p)$ such that for all circuits γ ,

$$p < 1/2 \text{ implies } c_1(p)|\partial\dot{\gamma}| \leq ES_\gamma, \text{Var } S_\gamma \leq c_2(p)|\partial\dot{\gamma}| \tag{1.1}$$

and

$$p > 1/2 \text{ implies } c_1(p)|\dot{\gamma}| \leq ES_\gamma, \text{Var } S_\gamma \leq c_2(p)|\dot{\gamma}| \tag{1.2}$$

Furthermore, if $\{\gamma(n)\}$ is any sequence of circuits with $|\dot{\gamma}(n)| \rightarrow \infty$, then $(S_{\gamma(n)} - ES_{\gamma(n)})/(\text{Var } S_{\gamma(n)})^{1/2}$ converges in distribution to the standard normal law.

The estimates on ES_γ and $\text{Var } S_\gamma$ indicate an essential qualitative difference in the behavior of S_γ for p above or below the critical value of $1/2$. By making “regularity” assumptions on the circuits $\{\gamma(n)\}$ it is possible to obtain more precise results. In particular, if γ_n is the boundary of the square $[0, n] \times [0, n]$, then

$$p < 1/2 \text{ implies } \begin{cases} ES_{\gamma(n)} = 4(n-1)\lambda + O(\ln n) \\ \text{Var } S_{\gamma(n)} = 4(n-1)\Lambda_1 + O(\ln n) \end{cases} \tag{1.3}$$

and

$$p > 1/2 \text{ implies } \begin{cases} ES_{\gamma(n)} = (n - 1)^2 p_\infty + 4(n - 1)\lambda + O(\ln n) \\ \text{Var } S_{\gamma(n)} = (n - 1)^2 \Lambda_2 + O(n \ln n) \end{cases} \quad (1.4)$$

where the constants $\lambda, p_\infty, \Lambda_1, \Lambda_2$ are defined in Section 5.

A more general central limit theorem which applies to sums of functions of the open clusters can be formulated as follows. A function f is said to be *increasing* (*decreasing*) on the subsets of \mathbb{Z}_2 if $f(W_1) \leq f(W_2)$ (\geq) whenever $W_1 \subset W_2$. Let \mathfrak{F} be the set of (finite) real valued functions f defined on the connected subsets of \mathbb{Z}_2 which are either increasing or decreasing, and are *constant* on infinite sets [i.e., $f(W_1) = f(W_2)$ if $|W_1| = |W_2| = +\infty$].

Theorem 2. Assume $p \neq 1/2$ and $\{\gamma(n)\}$ is a sequence of circuits with $|\dot{\gamma}(n)| \rightarrow \infty$. Assume $\{f_n\}$ is a sequence of functions satisfying the following:

$$f_n \in \mathfrak{F} \text{ for each } n \quad (1.5)$$

$$\sup_n \max_{x \in \dot{\gamma}(n)} E|f_n(W_x)|^k = C_k < \infty \text{ for each } k = 1, 2, \dots \quad (1.6)$$

$$\inf_n \min_{x \in \dot{\gamma}(n)} \text{Var}[f_n(W_x)] = \sigma^2 > 0 \quad (1.7)$$

If $S_{\gamma(n)} = \sum_{x \in \dot{\gamma}(n)} f_n(W_x)$, then $(S_{\gamma(n)} - ES_{\gamma(n)})/(\text{Var } S_{\gamma(n)})^{1/2}$ converges in distribution to the standard normal law.

Several remarks are in order here. It will be shown in Section 4 that Theorem 2 covers the $p > 1/2$ case of Theorem 1, but not the $p < 1/2$ case. The problem is condition (1.7). It will be seen in Section 2 [see (2.1)] that (1.6) is not overly restrictive. For the case $p > 1/2$, the number $S_\gamma = \sum_{x \in \dot{\gamma}} \xi_\gamma(x)$ of sites inside γ which are joined to γ by open paths may be divided into two components

$$S_\gamma = S_\gamma^I + S_\gamma^F$$

where

$$S_\gamma^I = \sum_{x \in \dot{\gamma}} 1_{\{|W_x| = +\infty\}}, S_\gamma^F = \sum_{x \in \dot{\gamma}} \eta_\gamma(x),$$

and $\eta_\gamma(x) = 1_{\{|W_x| < \infty, W_x \cap \gamma \neq \emptyset\}}$. The functions $f_n(W_x) = 1_{\{|W_x| = +\infty\}}$ satisfy the hypotheses of Theorem 2 for any sequence $\{\gamma(n)\}$, and we obtain a slightly more general version of a result of Grimmett (see Ref. 6), which is a central limit theorem for the number of sites inside $\gamma(n)$ which are "connected to infinity." The second component S_γ^F of S_γ is also asymptoti-

cally normally distributed; in fact it follows easily from Lemma 1 and the proof of (1.1) that there exist constants $c'_1(p)$ and $c'_2(p)$ such that

$$c'_1(p)|\partial\dot{\gamma}| \leq ES_{\dot{\gamma}}^F, \text{Var } S_{\dot{\gamma}}^F \leq c'_2(p)|\partial\dot{\gamma}|$$

and that $(S_{\dot{\gamma}(n)}^F - ES_{\dot{\gamma}(n)}^F)/(\text{Var } S_{\dot{\gamma}(n)}^F)^{1/2}$ converges to the standard normal law.

It is also possible to consider $f_n(W_x) = |W_{x,n}|^{-1} 1_{\{|W_{x,n}| > 0\}}$ where $W_{x,n}$ is the set of sites in $\dot{\gamma}(n)$ joined to x by open paths within $\dot{\gamma}(n)$. Although f_n is not monotone, it is still possible to prove asymptotic normality for $\sum_{x \in \dot{\gamma}(n)} f_n(W_x)$, the number of open clusters in $\dot{\gamma}(n)$.

It should be noted here that Theorem 2 above is similar to Theorem (3.1) in Neaderhouser's paper,⁽¹³⁾ except that strict regularity requirements for the circuits $\gamma(n)$ are imposed there. Although the approach used in Ref. 13 may possibly be modified to allow arbitrary circuits $\gamma(n)$, we feel that Malyšev's method of Ref. 12 using the method of moments and semi-invariants is simpler, and works easily for both Theorems 1 and 2. Malyšev's technique is also used in Refs. 1 and 9.

2. THE BASIC INEQUALITIES

Lemma 1. If $p \neq 1/2$ then there are finite nonzero constants α and β depending only on p such that

$$P(|W_0| < \infty \text{ and there exists } y \in W_0, d(0, y) \geq m) \leq \alpha e^{-\beta m}$$

Proof. For $p < 1/2$ W_0 is finite with probability 1, and part of Theorem 2 in Ref. 8 states that there exists some $\beta_1(p)$, $0 < \beta_1(p) < \infty$, such that $P(\text{there exists } y \in W_0, d(0, y) \geq m) \leq 2e^{-\beta_1(p)m}$. For $p > 1/2$ we turn to the dual-lattice technique, explained in Ref. 15. If $|W_0| < \infty$ and there exists $y \in W_0$ with $d(0, y) \geq m$, then there must exist some circuit of closed bonds in the dual-lattice containing W_0 (and hence the origin) with length at least m . Such a circuit must contain at least one of the dual-lattice sites $x_k^* = (k + 1/2, 1/2), k \geq 0$. Therefore, $P(|W_0| < \infty \text{ and there exists } y \in W_0, d(0, y) \geq m) \leq$

$$\begin{aligned} & \sum_{k=0}^{\infty} P(\text{there is a path in the dual-lattice containing } x_k^* \\ & \quad \text{with length } \geq \max(m, k)) \\ & \leq \sum_{k=0}^m 2e^{-\beta_1(1-p)m} + \sum_{k=m+1}^{\infty} 2e^{-\beta_1(1-p)k} \\ & \leq \alpha e^{-\beta m} \end{aligned}$$

for an appropriate choice of α, β depending only on p . ■

Corollary. If $p \neq 1/2$ and A is a finite subset of \mathbb{Z}_2 , then $P(|W_A| < \infty$ and there exists $y \in W_A$ with $d(x, y) \geq m$ for all $x \in A) \leq \alpha|A|e^{-\beta m}$.

Throughout the rest of the paper α and β will be the constants defined in Lemma 1. It follows from Lemma 1 that condition (1.6) of Theorem 2 is satisfied for a sequence $\{f_n\}$ if there is a function g such that $|g(\infty)| < \infty$, $\sup_n |f_n(W_x)| \leq g(|W_x|)$ and

$$\sum_{n=1}^{\infty} g(n)^k \exp(-\beta_1 n^{1/2}) < \infty \quad \text{for all } k = 1, 2, \dots \quad (2.1)$$

Remark. Since H. Kesten has recently shown (personal communication) that $P(|W_0| \geq n) \leq \alpha' e^{-\beta' n}$ for $p < 1/2$ where α', β' depend only on p , (2.1) can be improved accordingly.

We can now state and prove a result which will show that $f_n(W_x)$ is more or less “independent” of $f_n(W_y)$ if x and y are “far apart.” Let $d(x, A) = \min\{d(x, y) \mid y \in A\}$ and $d(A, B) = \min\{d(x, B) \mid x \in A\}$.

Lemma 2. Assume $p \neq 1/2$ and $\{\gamma_n\}$ is a sequence of circuits with $|\dot{\gamma}(n)| \rightarrow \infty$. Let $\{f_n\}$ be a sequence of functions which satisfy conditions (1.5) and (1.6) and the additional requirement that $f_n(W_x) = 0$ if $|W_x| = +\infty$. For finite sets $A \subset \mathbb{Z}_2$ let $\rho_n(A) = \prod_{x \in A} f_n(W_x)$. Then for all finite sets $A, B \subset \mathbb{Z}_2$ there exists a finite constant c_3 depending only on $p, |A|, |B|$, and the numbers C_k in (1.6) such that for all n ,

$$|E\rho_n(A)\rho_n(B) - E\rho_n(A)E\rho_n(B)| \leq c_3 e^{-\beta d(A,B)/4}$$

Proof. We first observe that repeated application of Hölder’s inequality to $E|\rho_n(A)|$ and condition (1.6) show that $E|\rho_n(A)|$ is bounded above by a quantity which depends only on $p, |A|$, and the numbers C_k in (1.6). Let $m = d(A, B)$ and let $\Omega_A = \{d(x, A) \leq m/2 \text{ for all } x \in W_A\}$ and $\Omega_B = \{d(x, B) \leq m/2 \text{ for all } x \in W_B\}$. Then Ω_A and Ω_B are independent, and writing $E(X; G)$ for $E(X \mid G)$, $E(\rho_n(A)\rho_n(B); \Omega_A \cap \Omega_B) = E(\rho_n(A); \Omega_A)E(\rho_n(B); \Omega_B)$. Thus

$$\begin{aligned} |E\rho_n(A)\rho_n(B) - E\rho_n(A)E\rho_n(B)| &\leq |E(\rho_n(A)\rho_n(B); \Omega_A^c \cup \Omega_B^c)| \\ &\quad + |E(\rho_n(A); \Omega_A)E(\rho_n(B); \Omega_B) - E\rho_n(A)E\rho_n(B)| \\ &\leq E(|\rho_n(A)\rho_n(B)|; \Omega_A^c \cup \Omega_B^c) + E|\rho_n(A)|E(|\rho_n(B)|; \Omega_B^c) \\ &\quad + E|\rho_n(B)|E(|\rho_n(A)|; \Omega_A^c) \end{aligned}$$

Since $\rho_n(A) = 0$ if $|W_A| = +\infty$ and $\rho_n(B) = 0$ if $|W_B| = +\infty$, we can let $\tilde{\Omega}_A = \Omega_A^c \cap \{|W_A| < \infty\}$ and $\tilde{\Omega}_B = \Omega_B^c \cap \{|W_B| < \infty\}$ and bound the terms

above by

$$E(|\rho_n(A)\rho_n(B)|; \tilde{\Omega}_A) + E(|\rho_n(A)\rho_n(B)|; \tilde{\Omega}_B) + E|\rho_n(A)|E(|\rho_n(B)|; \tilde{\Omega}_B) + E|\rho_n(B)|E(|\rho_n(A)|; \tilde{\Omega}_A)$$

Replacing $\tilde{\Omega}_A$ and $\tilde{\Omega}_B$ with their indicator functions and using Hölder's inequality we have

$$\begin{aligned} &|E\rho_n(A)\rho_n(B) - E\rho_n(A)E\rho_n(B)| \\ &\leq [E\rho_n^2(A)\rho_n^2(B)]^{1/2} [(E1_{\tilde{\Omega}_A})^{1/2} + (E1_{\tilde{\Omega}_B})^{1/2}] \\ &\quad + E|\rho_n(A)| [E\rho_n^2(B)]^{1/2} (E1_{\tilde{\Omega}_B})^{1/2} \\ &\quad + E|\rho_n(B)| [E\rho_n^2(A)]^{1/2} (E1_{\tilde{\Omega}_A})^{1/2} \\ &\leq c_3 e^{-\beta m/4} \end{aligned}$$

using the Corollary to Lemma 1 to estimate $E1_{\tilde{\Omega}_A}$ and $E1_{\tilde{\Omega}_B}$, where c_3 depends on $p, |A|, |B|$, and the C_k in (1.6). ■

3. PROOF OF THEOREM 2

We start by pointing out that it suffices to consider sequences $\{f_n\}$ such that $f_n(W_x) = 0$ if $|W_x| = +\infty$. This is because any $f_n \in \mathcal{F}$ can be written as $f_n = d_n + g_n$, where $d_n = f_n(\mathbb{Z}_2)$ and $g_n(W_x) = [f_n(W_x) - d_n] 1_{\{|W_x| < \infty\}}$. Next we observe that the FKG inequalities (see Refs. 3, 4, and 16) imply that $Ef_n(W_x)f_n(W_y) - Ef_n(W_x)Ef_n(W_y) \geq 0$, so that

$$\begin{aligned} \text{Var } S_{\gamma(n)} &= \sum_{x \in \tilde{\gamma}(n)} \sum_{y \in \tilde{\gamma}(n)} Ef_n(W_x)f_n(W_y) - Ef_n(W_x)Ef_n(W_y) \\ &\geq \sum_{x \in \tilde{\gamma}(n)} \text{Var } f_n(W_x) \\ &\geq \sigma^2 |\tilde{\gamma}(n)| \end{aligned}$$

by condition (1.7).

It is now necessary to recall certain facts about semi-invariants and Ursell functions, which can be found in Refs. 9, 11, and 12. The k th semi-invariant of a random variable X will be written $\nu_k(X)$; it is the coefficient of t^k in the expansion of $\ln Ee^{tX}$, and can be expressed in terms of the moments of X by

$$\nu_k(X) = \sum_{m=1}^k (-1)^{m-1} \frac{1}{m} \sum_{\substack{r_1, r_2, \dots, r_m \geq 1 \\ r_1 + r_2 + \dots + r_m = k}} \frac{k!}{r_1! r_2! \dots r_m!} EX^{r_1} EX^{r_2} \dots EX^{r_m} \tag{3.1}$$

The Ursell function $\nu(X_1, X_2, \dots, X_n)$ of n variables is

$$\nu(X_1, \dots, X_n) = \sum_{s=1}^n (-1)^{s-1} (s-1)! \sum_{\underline{\pi} : |\underline{\pi}|=s} E(X_{i_1} X_{i_2} \dots) \times E(X_{i_1^2} X_{i_2^2} \dots) \cdots E(X_{i_1^s} X_{i_2^s} \dots) \tag{3.2}$$

where the second sum is over all partitions $\underline{\pi} = \{\{i_1^1, i_2^1, \dots\}, \{i_1^2, i_2^2, \dots\} \cdots \{i_1^s, i_2^s, \dots\}\}$ of $\{1, 2, \dots, n\}$ consisting of s members. It is known (see Refs. 9, 11, 12) that the semi-invariants of a sum can be written as

$$\nu_k \left(\sum_{i=1}^n X_i \right) = \sum_{i_1=1}^n \cdots \sum_{i_k=1}^n \nu(X_{i_1}, \dots, X_{i_k}) \tag{3.3}$$

Since the moments are determined by the semi-invariants, to prove $Z_n = (S_{\gamma(n)} - ES_{\gamma(n)}) / (\text{Var } S_{\gamma(n)})^{1/2}$ converges in distribution to the standard normal law it suffices to prove $\nu_k(Z_n) \rightarrow 0$ if $k \neq 2$ and $\nu_2(Z_n) \rightarrow 1$ (the semi-invariants of the standard normal law). Note that $\nu_1(Z_n) = EZ_n = 0$ and $\nu_2(Z_n) = \text{Var } Z_n = 1$.

For any finite set $A \subset \mathbb{Z}_2$ define $G_{k,m}(A)$ by

$$G_{k,m}(A) = \left\{ (x_1, x_2, \dots, x_k) \mid \text{each } x_i \in A \text{ and the maximum over nontrivial partitions } \{\pi', \pi''\} \text{ of } \{1, 2, \dots, k\} \text{ of } \min_{i \in \pi', j \in \pi''} d(x_i, x_j) \text{ is equal to } m \right\} \tag{3.4}$$

A counting argument (see Ref. 9) shows that

$$|G_{k,m}(A)| \leq |A|(2m+1)^{2k} (k!)^2 \tag{3.5}$$

The last formula needed (see Refs. 9, 11, 12) expresses the Ursell functions in terms of moments. Let $\{\pi', \pi''\}$ be a nontrivial partition of $\{1, 2, \dots, k\}$, fix $(x_1, x_2, \dots, x_k) \in G_{k,m}$, and let $\rho_n(\pi) = \prod_{i \in \pi} f_n(W_{x_i})$ for any subset $\pi \subset \{1, 2, \dots, k\}$. Then

$$\begin{aligned} &\nu(f_n(W_{x_1}), f_n(W_{x_2}), \dots, f_n(W_{x_k})) \\ &= \sum_{\substack{\underline{\pi} = \{\pi_1, \pi_2, \pi_3, \dots\} \\ \pi_1 \subset \pi', \pi_2 \subset \pi''}} \pm [E\rho_n(\pi_1)\rho_n(\pi_2) \\ &\quad - E\rho_n(\pi_1)E\rho_n(\pi_2)] E\rho_n(\pi_3)E\rho_n(\pi_4) \cdots \end{aligned} \tag{3.6}$$

where the sum is over nontrivial partitions $\underline{\pi}$ of $\{1, 2, \dots, k\}$ and the sign depends on $\underline{\pi}$. For each $(x_1, x_2, \dots, x_k) \in G_{k,m}$, by choosing the particular

$\{\pi', \pi''\}$ such that $\min_{i \in \pi', j \in \pi''} d(x_i, x_j) = m$, and using Lemma 2, we obtain.

$$|\nu(f_n(W_{x_1}), \dots, f_n(W_{x_k}))| \leq c'_3 \cdot \Gamma_k \cdot e^{-\beta m/4} \tag{3.7}$$

where c'_3 is $\sup_n \max_{\pi \subset \{1, 2, \dots, k\}} [1 + E\rho_n^2(\pi)]^k$ and Γ_k is the number of partitions of $\{1, 2, \dots, k\}$.

Using (3.7) in (3.3) we obtain

$$\begin{aligned} \nu_k(Z_n) &\leq (\text{Var } S_{\dot{\gamma}(n)})^{-k/2} \sum_{m=0}^{\infty} \sum_{(x_1, \dots, x_k) \in G_{k,m}(\dot{\gamma}(n))} \nu(f_n(W_{x_1}), \dots, f_n(W_{x_k})) \\ &\leq \sigma^{-k} |\dot{\gamma}(n)|^{-k/2} \sum_{m=0}^{\infty} |\dot{\gamma}(n)| (2m+1)^k (k!)^2 c'_3 \Gamma_k e^{-\beta m/4} \\ &= |\dot{\gamma}(n)|^{1-k/2} \sigma^{-k} c'_3 (k!)^2 \Gamma_k \sum_{m=0}^{\infty} (2m+1)^k e^{-\beta m/4} \\ &\rightarrow 0 \end{aligned}$$

for $k \geq 3$.

4. PROOF OF THEOREM 1

We will consider only the case $p < 1/2$, since for $p > 1/2$,

$$\begin{aligned} \text{Var } \xi_{\dot{\gamma}}(x) &= P(W_x \cap \gamma \neq \emptyset) P(W_x \cap \gamma = \emptyset) \\ &\geq P(|W_x| = +\infty) \cdot (1-p)^4 \end{aligned}$$

a bound which implies Theorem 2 applies. Throughout the remainder of this section we assume $p < 1/2$. We will first prove the variance estimates of (1.1), omitting the similar estimates on $ES_{\dot{\gamma}}$. Since the FKG inequality implies $E\xi_{\dot{\gamma}}(x)\xi_{\dot{\gamma}}(y) - E\xi_{\dot{\gamma}}(x)E\xi_{\dot{\gamma}}(y) \geq 0$,

$$\begin{aligned} \text{Var } S_{\dot{\gamma}} &= \sum_{x \in \dot{\gamma}} \sum_{y \in \dot{\gamma}} E\xi_{\dot{\gamma}}(x)\xi_{\dot{\gamma}}(y) - E\xi_{\dot{\gamma}}(x)E\xi_{\dot{\gamma}}(y) \\ &\geq \sum_{x \in \partial \dot{\gamma}} E\xi_{\dot{\gamma}}(x) [1 - E\xi_{\dot{\gamma}}(x)] \\ &= \sum_{x \in \partial \dot{\gamma}} P(W_x \cap \gamma \neq \emptyset) P(W_x \cap \gamma = \emptyset) \end{aligned}$$

Since $x \in \partial \dot{\gamma}$, $P(W_x \cap \gamma \neq \emptyset) \geq p$ and $P(W_x \cap \gamma = \emptyset) \geq (1-p)^4$, and $\text{Var } S_{\dot{\gamma}} \geq p(1-p)^4 |\partial \dot{\gamma}|$. To obtain an upper bound we introduce $R_i(\gamma) = \{x \in \dot{\gamma} \mid d(x, \partial \dot{\gamma}) = i\}$ and $\tilde{\xi}_{\dot{\gamma}}(x) = 1 - \xi_{\dot{\gamma}}(x)$. Note that $E\xi_{\dot{\gamma}}(x)\xi_{\dot{\gamma}}(y) - E\xi_{\dot{\gamma}}(x)E\xi_{\dot{\gamma}}(y) = E\tilde{\xi}_{\dot{\gamma}}(x)\tilde{\xi}_{\dot{\gamma}}(y) - E\tilde{\xi}_{\dot{\gamma}}(x)E\tilde{\xi}_{\dot{\gamma}}(y)$. If we let $i_{\dot{\gamma}} = \max\{i \mid R_i(\gamma) \neq \emptyset\}$, it

is clear that $i_\gamma \leq |\partial\dot{\gamma}|$.

$$\begin{aligned} \text{Var } S_\gamma &= \sum_{i=1}^{i_\gamma} \sum_{x \in R_i(\gamma)} \sum_{y \in \dot{\gamma}} E \xi_\gamma(x) \xi_\gamma(y) - E \xi_\gamma(x) E \xi_\gamma(y) \\ &= \sum_{i=1}^{i_\gamma} \sum_{x \in R_i(\gamma)} \sum_{\substack{y \in \dot{\gamma} \\ d(x,y) \leq i}} E \xi_\gamma(x) \xi_\gamma(y) - E \xi_\gamma(x) E \xi_\gamma(y) \\ &\quad + \sum_{i=1}^{i_\gamma} \sum_{x \in R_i(\gamma)} \sum_{\substack{y \in \dot{\gamma} \\ d(x,y) > i}} E \tilde{\xi}_\gamma(x) \tilde{\xi}_\gamma(y) - E \tilde{\xi}_\gamma(x) E \tilde{\xi}_\gamma(y) \\ &\leq \sum_{i=1}^{\infty} \sum_{x \in R_i(\gamma)} \sum_{\substack{y \in \dot{\gamma} \\ d(x,y) \leq i}} P(W_x \cap \gamma \neq \emptyset) \\ &\quad + \sum_{i=1}^{\infty} \sum_{x \in R_i(\gamma)} \sum_{\substack{y \in \dot{\gamma} \\ d(x,y) > i}} c_3 e^{-\beta d(x,y)/4} \end{aligned}$$

using Lemma 2, which applies because $\tilde{\xi}_\gamma(x) = 0$ if $|W_x| = +\infty$. We can now use the rather crude estimate $|R_i(\gamma)| \leq 4\sqrt{2} i |\partial\dot{\gamma}|$, for $i \geq 1$, and Lemma 1 to bound the preceding terms by

$$4\sqrt{2} |\partial\dot{\gamma}| \left\{ \sum_{i=1}^{\infty} i(+2i+1)^2 \alpha e^{-\beta i} + \sum_{i=1}^{\infty} \sum_{\substack{y \in \mathbb{Z}_2 \\ d(0,y) \geq i}} i c_3 e^{-\beta d(0,y)/4} \right\}$$

which proves (1.1).

To prove $S_{\gamma(n)}$ is approximately normal we will use the method of Section 3 after we show that only the terms $\xi_{\gamma(n)}(x)$ with x near $\partial\dot{\gamma}(n)$ contribute significantly to $S_{\gamma(n)}$. To do this, define $T_K(\gamma) = \{x \in \dot{\gamma} \mid d(x, \partial\dot{\gamma}) \geq K\}$, where K will be chosen later. The terms of $S_{\gamma(n)}$ which come from $T_K(\gamma(n))$ have variance

$$\begin{aligned} &\text{Var} \left(\sum_{x \in T_K(\gamma(n))} \xi_\gamma(x) \right) \\ &= \sum_{x \in T_K(\gamma(n))} \sum_{y \in T_K(\gamma(n))} E \xi_{\gamma(n)}(x) \xi_{\gamma(n)}(y) - E \xi_{\gamma(n)}(x) E \xi_{\gamma(n)}(y) \\ &\leq \sum_{x \in T_K(\gamma(n))} \sum_{y \in T_K(\gamma(n))} P(\text{there exists } z \in W_x, d(z, x) \geq K) \\ &\leq |T_K(\gamma(n))|^2 \alpha e^{-\beta K} \end{aligned}$$

by Lemma 1. The additional crude estimates of $|T_K(\gamma(n))| \leq |\dot{\gamma}(n)|^2$ and $|\dot{\gamma}(n)| = \sum_{i=1}^K |R_i(\gamma(n))| \leq 4\sqrt{2} |\partial\dot{\gamma}(n)|^3$, with the choice of $K = |\partial\dot{\gamma}(n)|^{1/5}$, give

$$\text{Var}\left(\sum_{x \in T_K(\gamma(n))} \xi_{\gamma(n)}(x)\right) \rightarrow 0$$

It suffices then to check $\nu_k(\sum_{x \in \dot{\gamma}(n) \setminus T_K(\gamma(n))} \xi_{\gamma(n)}(x) / (\text{Var } S_{\gamma(n)})^{1/2}) \rightarrow 0$ for $k \geq 3$. As in Section 3 we bound ν_k by

$$\begin{aligned} & (\text{Var } S_{\gamma(n)})^{-k/2} \sum_{m=0}^{\infty} |\dot{\gamma}(n) \setminus T_K(\gamma(n))| (2m+1)^k (k!)^2 \Gamma_k c_3 e^{-\beta m/4} \\ & \leq (p(1-p)^4)^{-k/2} |\partial\dot{\gamma}(n)|^{-k/2} |\dot{\gamma}(n) \setminus T_K(\gamma(n))| \cdot c_3 \Gamma_k (k!)^2 \\ & \sum_{m=0}^{\infty} (2m+1)^k e^{-\beta m/4} \rightarrow 0 \end{aligned}$$

for $k \geq 3$, since with $K = |\partial\dot{\gamma}(n)|^{1/5}$,

$$\begin{aligned} |\dot{\gamma}(n) \setminus T_K(\dot{\gamma}(n))| & \leq \sum_{i=0}^K |R_i(\gamma(n))| \leq |\partial\dot{\gamma}(n)| \left(1 + \frac{4\sqrt{2} K(K+1)}{2}\right) \\ & = O(|\partial\dot{\gamma}(n)|^{1+2/5}) \quad \blacksquare \end{aligned}$$

5. ESTIMATES FOR THE SQUARES

In this section $\gamma(n)$ will be the boundary of the square $[0, n] \times [0, n]$, and we will write S_n for $S_{\gamma(n)}$ and $\xi_n(x)$ for $\xi_{\gamma(n)}(x)$. The constants in (1.3) and (1.4) are

$$\begin{aligned} p_{\infty} &= P(|W_0| = +\infty) \\ \lambda &= \sum_{n=1}^{\infty} P(|W_0| < \infty \text{ and } W_0 \cap H_n \neq \emptyset) \\ \Lambda_2 &= \sum_{y \in \mathbb{Z}_2} P(|W_0| = |W_y| = +\infty) - p_{\infty}^2 \\ \Lambda_1 &= \sum_{n=1}^{\infty} \sum_{\substack{y \in \mathbb{Z}_2 \\ y_1 < n}} \{P(W_0 \cap H_n \neq \emptyset \text{ and } W_y \cap H_n \neq \emptyset) \\ & \quad - P(W_0 \cap H_n \neq \emptyset) \cdot P(W_y \cap H_n \neq \emptyset)\} \end{aligned} \tag{5.1}$$

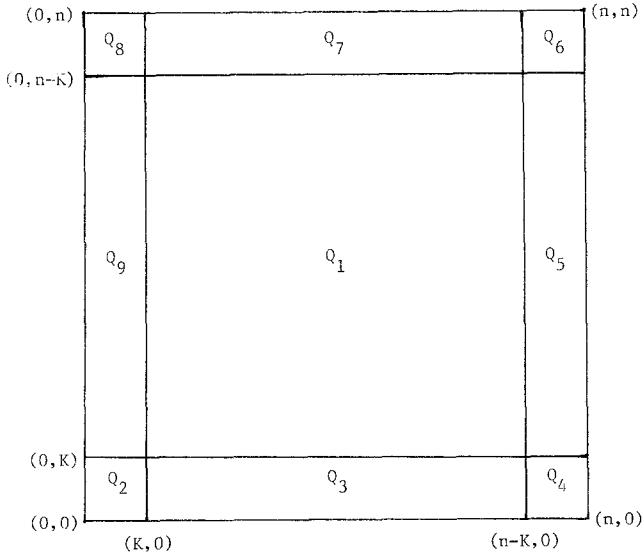


Fig. 1

where $H_n = \{x \in \mathbb{Z}_2 \mid x_1 = n\}$. The fact that Λ_1 and Λ_2 are finite follows from Lemmas 1 and 2. We will prove the estimates in (1.3) and omit the similar proof of (1.4).

It is convenient to divide the square as follows (where K is a number to be chosen later) (Fig. 1):

i.e., $Q_1 = \{x \mid K \leq x_1 \leq n - K, K \leq x_2 \leq n - K\}$, $Q_2 = \{x \mid 0 < x_1 < K, 0 < x_2 < K\}$, $Q_3 = \{x \mid K \leq x_1 \leq n - K, 0 < x_2 < K\}$, and so on. Note that $ES_n = \sum_{i=1}^9 \sum_{x \in Q_i} P(W_x \cap \gamma_n \neq \emptyset)$. We start with Q_1 :

$$\sum_{x \in Q_1} P(W_x \cap \gamma(n) \neq \emptyset) \leq n^2 \alpha e^{-\beta K} \tag{5.2}$$

For $1 \leq i \leq K - 1$ define $l_i = \{x \in Q_2 \mid \min(x_1, x_2) = i\}$. Then

$$\begin{aligned} \sum_{x \in Q_2} P(W_x \cap \gamma(n) \neq \emptyset) &= \sum_{i=1}^{K-1} \sum_{x \in l_i} P(W_x \cap \gamma(n) \neq \emptyset) \\ &\leq \sum_{i=1}^{K-1} |l_i| P(\text{there exists } y \in W_x, d(x, y) \geq i) \\ &\leq 2\alpha K \sum_{i=1}^{\infty} e^{-\beta i} \end{aligned} \tag{5.3}$$

using Lemma 1. Now let l_0^* be the x axis and for $1 \leq i \leq K-1$ define $l_i^* = \{x \in Q_3 \mid x_2 = i\}$. Then

$$\begin{aligned} \sum_{x \in Q_3} P(W_x \cap \gamma(n) \neq \emptyset) &= \sum_{i=1}^{K-1} \sum_{x \in l_i^*} \{P(W_x \cap l_0^* \neq \emptyset) \\ &\quad + P(W_x \cap l_0^* = \emptyset, W_x \cap \gamma(n) \neq \emptyset)\} \\ &= \sum_{i=1}^{K-1} (n+1-2K)P(W_0 \cap H_i \neq \emptyset) \tag{5.4} \\ &\quad + \sum_{i=1}^{K-1} \sum_{x \in l_i^*} P(W_x \cap \gamma(n) \neq \emptyset, W_x \cap l_0^* = \emptyset) \end{aligned}$$

It is not difficult to show that the first term above is $(n-1)\lambda$ plus a term bounded by

$$\begin{aligned} 2K\lambda + n \sum_{i=K}^{\infty} P(W_0 \cap H_i \neq \emptyset) &\leq 2K\lambda + n \sum_{i=K}^{\infty} \alpha e^{-\beta i} \\ &= 2K\lambda + n\alpha e^{-\beta K} / (1 - e^{-\beta}) \end{aligned}$$

The second term is bounded by

$$KnP(\text{there exists } x \in W_0 \text{ with } d(0, x) \geq K) \leq Kn\alpha e^{-\beta K}$$

The terms in Q_4 through Q_9 are similar, and therefore

$$\begin{aligned} |ES_n - 4(n-1)\lambda| &\leq n^2\alpha e^{-\beta K} + 8\alpha K \sum_{i=1}^{\infty} e^{-\beta i} + 8K\lambda + 4n\alpha e^{-\beta K} / (1 - e^{-\beta}) \\ &\quad + 4Kn\alpha e^{-\beta K} \end{aligned}$$

The choice $K = 2\beta^{-1} \ln n$ yields $ES_n = 4(n-1)\lambda + O(\ln n)$. We turn now to the variance estimate

$$\begin{aligned} \text{Var } S_n &= \sum_{x \in \tilde{\gamma}(n)} \left(\sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} + \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) > K}} \right) \{E\xi_n(x)\xi_n(y) - E\xi_n(x)E\xi_n(y)\} \\ &= \sum_{x \in \tilde{\gamma}(n)} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} (E\xi_n(x)\xi_n(y) - E\xi_n(x)E\xi_n(y) + O(n^4 e^{-\beta K/4})) \end{aligned}$$

using Lemma 2 and the fact that $\xi_{\tilde{\gamma}(n)}$ can be replaced with $\tilde{\xi}_{\tilde{\gamma}(n)}$. We now use the decomposition of $\gamma(n)$.

$$\sum_{x \in Q_1} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} E\xi_n(x)\xi_n(y) - E\xi_n(x)E\xi_n(y) \leq n^2(2K+1)^2\alpha e^{-\beta K} \tag{5.5}$$

by Lemma 1.

$$\begin{aligned}
 & \sum_{x \in Q_2} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} E \xi_n(x) \xi_n(y) - E \xi_n(x) E \xi_n(y) \\
 &= \sum_{i=1}^{K-1} \sum_{x \in I_i} \left(\sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq i}} + \sum_{\substack{y \in \tilde{\gamma}(n) \\ i < d(x,y) \leq K}} \right) \\
 & \quad \cdot \{ E \xi_n(x) \xi_n(y) - E \xi_n(x) E \xi_n(y) \} \\
 & \leq \sum_{i=1}^{\infty} 2K \left\{ (2i+1)^2 \cdot 2\alpha e^{-\beta i} + \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \geq i}} c_3 e^{-\beta d(x,y)/4} \right\} \\
 &= O(K) \tag{5.6}
 \end{aligned}$$

$$\begin{aligned}
 & \sum_{x \in Q_3} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} E \xi_n(x) \xi_n(y) - E \xi_n(x) E \xi_n(y) \\
 &= \sum_{x \in Q_3} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} P(W_x \cap I_0^* \neq \emptyset, W_y \cap I_0^* \neq \emptyset) \\
 & \quad - P(W_x \cap I_0^* \neq \emptyset) P(W_y \cap I_0^* \neq \emptyset) \\
 & \quad + \sum_{x \in Q_3} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} \{ P(W_x \cap \gamma(n) \neq \emptyset, \\
 & \quad W_y \cap \gamma(n) \neq \emptyset, W_x \cap I_0^* = \emptyset \text{ or } W_y \cap I_0^* = \emptyset) \\
 & \quad - P(W_x \cap I_0^* \neq \emptyset) \cdot P(W_y \cap I_0^* = \emptyset, W_y \cap \gamma(n) \neq \emptyset) \\
 & \quad - P(W_y \cap \gamma(n) \neq \emptyset) P(W_x \cap I_0^* = \emptyset, W_x \cap \gamma(n) \neq \emptyset) \} \tag{5.7}
 \end{aligned}$$

Since $P(W_x \cap I_0^* = \emptyset, W_x \cap \gamma(n) \neq \emptyset) \leq P(\text{there exists } y \in W_x, d(x, y) \geq K) \leq \alpha e^{-\beta K}$, we have

$$\begin{aligned}
 & \left| \sum_{x \in Q_3} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} (E \xi_n(x) \xi_n(y) - E \xi_n(x) E \xi_n(y)) \right. \\
 & \quad - \sum_{x \in Q_3} \sum_{\substack{y \in \tilde{\gamma}(n) \\ d(x,y) \leq K}} (P(W_x \cap I_0^* \neq \emptyset, W_y \cap I_0^* \neq \emptyset) \\
 & \quad \left. - P(W_x \cap I_0^* \neq \emptyset) P(W_y \cap I_0^* \neq \emptyset)) \right| \\
 & \leq 4(2K+1)^2 \cdot K \cdot \alpha \alpha e^{-\beta K}
 \end{aligned}$$

Now we can rewrite

$$\begin{aligned}
 & \sum_{x \in Q_3} \sum_{\substack{y \in \gamma(n) \\ d(x,y) \leq K}} P(W_x \cap I_0^* \neq \emptyset, W_y \cap I_0^* \neq \emptyset) \\
 & \quad - P(W_x \cap I_0^* \neq \emptyset)P(W_y \cap I_0^* \neq \emptyset) \\
 & = \sum_{x \in Q_3} \sum_{\substack{y \in \mathbb{Z}_2 \\ y_2 > 0}} P(W_x \cap I_0^* \neq \emptyset, W_y \cap I_0^* \neq \emptyset) \\
 & \quad - P(W_x \cap I_0^* \neq \emptyset)P(W_y \cap I_0^* \neq \emptyset) \\
 & \quad - \sum_{x \in Q_3} \sum_{\substack{y \in \mathbb{Z}_2 \\ y_2 > 0 \\ d(x,y) > K}} P(W_x \cap I_0^* \neq \emptyset, W_y \cap I_0^* \neq \emptyset) \\
 & \quad - P(W_x \cap I_0^* \neq \emptyset)P(W_y \cap I_0^* \neq \emptyset)
 \end{aligned}$$

The first term above is $(n-1)\Lambda_1 - 2(K-1)\Lambda_1$ and it is not difficult to show that the second term is at most $O(nKe^{-\beta K/4})$. Considering the terms from Q_4 to Q_9 we have

$$\begin{aligned}
 |\text{Var } S_n - 4(n-1)\Lambda_1| & \leq O(n^4 e^{-\beta K/4}) + 16n^2(2K+1)^2 \alpha e^{-\beta K} + O(K) \\
 & \quad + 2(K-1)\Lambda_1 + O(nKe^{-\beta K/4}) \\
 & = O(\ln n)
 \end{aligned}$$

for the choice $K = 16\beta^{-1} \ln n$. ■

NOTE ADDED IN PROOF

We have learned that Gunnar Branvall has independently obtained several central limit theorems similar to ours. He uses different techniques and works with circuits $\gamma(n)$ which are “regular.”

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